

**CEH**  
CES Engineering

# SIMULATION SYSTEMS

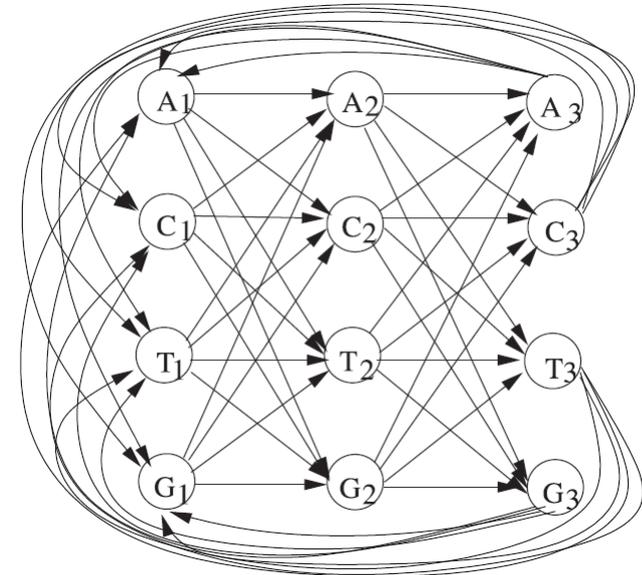
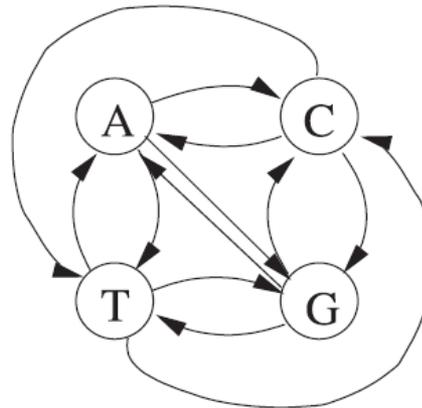
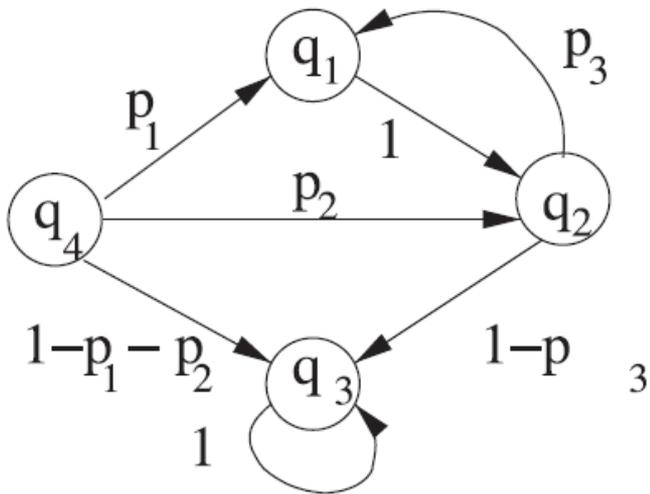
## MARKOV MODELS

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# Definition

2

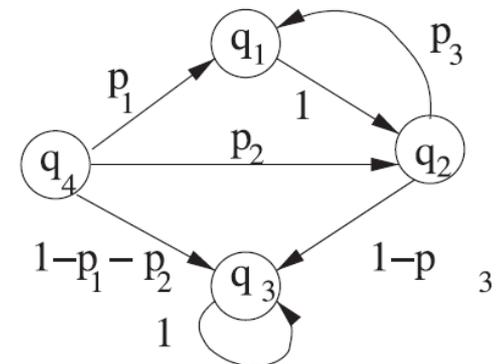
- A Markov model is generally represented as a graph containing a set of states represented as nodes and a set of transitions with probabilities represented by weighted edges.



# Simulation of Markov Models

3

- We simulate a Markov model by starting at some state and moving to successive neighboring states by choosing randomly among neighbors according to their labeled probabilities.
- For example, if we start in state  $q_4$ , then we would have probability  $p_1$  of moving to  $q_1$ ,  $p_2$  of moving to  $q_2$ , and  $(1-p_1-p_2)$  of moving to  $q_3$ . If we move to  $q_2$ , then we have probability  $p_3$  of moving to  $q_1$  and  $(1-p_3)$  of moving to  $q_3$ , and so on. The result is a walk through the state set (e.g.,  $q_1; q_2; q_1; q_2; q_3; q_3; \dots$ ).
- Resulting sequence of states is called a **“Markov chain”**



# Markov Model Components

4

- A state set  $Q = \{q_1; q_2; \dots; q_n\}$
- A starting distribution  $\Pr\{q(0) = q_i\} = p_i$ 
  - ▣ Represented by a vector  $\vec{p}$
- A set of transition probabilities:  
$$\Pr\{q(n+1) = q_j \mid q(n) = q_i\} = p_{ij}$$
  - ▣ Represented by a matrix  $\mathbf{P}$

This is the definition of the **First Order** Markov Model: probability of entering each possible next state dependent only on the current state

# Higher Order Markov Models

5

- $k^{\text{th}}$  Order Markov Model:

$$\Pr\{q(n) = q_{i,n} \mid q(n-1) = q_{i,(n-1)}, q(n-2) = q_{i,(n-2)}, \dots, q(n-k) = q_{i,(n-k)}\} = p_{i,j}$$

- ▣ Probability of next state depends on previous  $k$  states
- Note: Any  $k^{\text{th}}$ -order Markov model can be transformed into a first order Markov model by defining a new state set  $Q' = Q^k$  (i.e., each state in  $Q'$  is a set of  $k$  states in  $Q$ ), with current state in  $Q'$  being the last  $k$  states visited in  $Q$ .
  - ▣ Then a Markov chain in the  $k^{\text{th}}$ -order model  $Q = q_1; q_2; q_3; q_4; \dots$  —becomes the chain  $\{q_1; q_2; \dots; q_k\}; \{q_2; q_3; \dots; q_k\}; \{q_3; q_4; \dots; q_k\}; \dots$  in  $Q'$
  - ▣ Ignore higher-order Markov models when talking about theory

# Time Evolution of Markov Models

6

- Although the behavior of Markov models is random, it is also in some ways predictable
- Suppose we have a two-state model:  $Q=\{q_1; q_2\}$ , with initial probabilities  $p_1$  and  $p_2$  and transition probabilities  $p_{11}$ ,  $p_{12}$ ,  $p_{21}$ , and  $p_{22}$

- ▣ Step 0: 
$$\begin{bmatrix} Pr\{q(0) = q_1\} \\ Pr\{q(0) = q_2\} \end{bmatrix} = \begin{bmatrix} p_1 \\ p_2 \end{bmatrix}$$

- ▣ After 1 step: 
$$\begin{bmatrix} Pr\{q(1) = q_1\} \\ Pr\{q(1) = q_2\} \end{bmatrix} = \begin{bmatrix} p_1 p_{11} + p_2 p_{21} \\ p_1 p_{12} + p_2 p_{22} \end{bmatrix}$$

$$\begin{bmatrix} Pr\{q(2) = q_1\} \\ Pr\{q(2) = q_2\} \end{bmatrix} = \begin{bmatrix} (p_1 p_{11} + p_2 p_{21}) p_{11} + (p_1 p_{12} + p_2 p_{22}) p_{21} \\ (p_1 p_{11} + p_2 p_{21}) p_{12} + (p_1 p_{12} + p_2 p_{22}) p_{22} \end{bmatrix}$$

# Time Evolution of Markov Models

7

- Matrix Notation:

$$\begin{bmatrix} \Pr\{q(i+1) = q_1\} \\ \Pr\{q(i+1) = q_2\} \end{bmatrix} = \begin{bmatrix} p_{11} & p_{21} \\ p_{12} & p_{22} \end{bmatrix} \begin{bmatrix} \Pr\{q(i) = q_1\} \\ \Pr\{q(i) = q_2\} \end{bmatrix}$$

- Distribution after n steps:

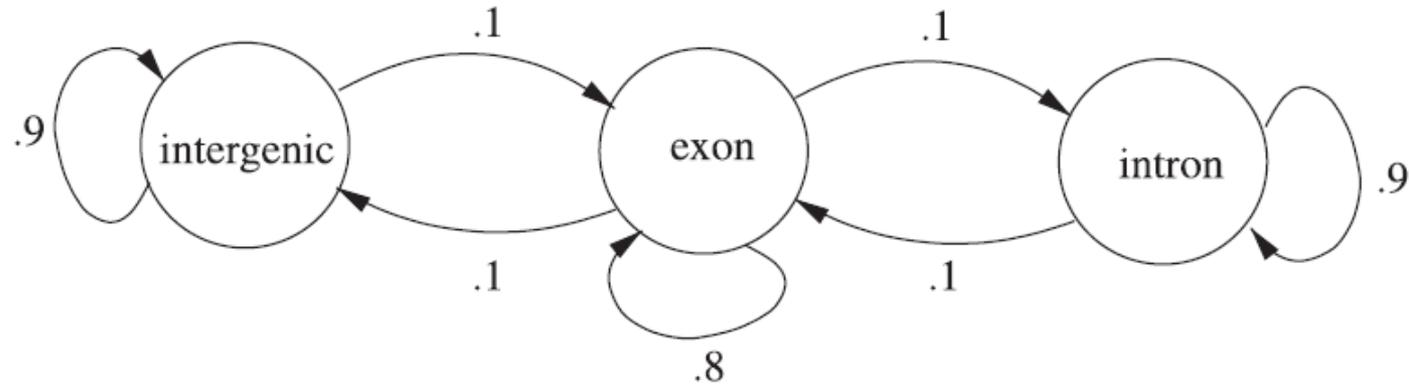
$$\begin{bmatrix} \Pr\{q(n) = q_1\} \\ \Pr\{q(n) = q_2\} \end{bmatrix} = \begin{bmatrix} p_{11} & p_{21} \\ p_{12} & p_{22} \end{bmatrix} \times \cdots \times \begin{bmatrix} p_{11} & p_{21} \\ p_{12} & p_{22} \end{bmatrix} \begin{bmatrix} p_1 \\ p_2 \end{bmatrix}$$

# Time Evolution: Example

8

$$P = \begin{bmatrix} .8 & .1 & .1 \\ .1 & .9 & 0 \\ .1 & 0 & .9 \end{bmatrix}$$

$$\begin{bmatrix} p_1 \\ p_2 \\ p_3 \end{bmatrix} = \begin{bmatrix} 0.2 \\ 0.3 \\ 0.5 \end{bmatrix}$$



$$\text{Time 0 } \begin{bmatrix} .2 \\ .3 \\ .5 \end{bmatrix} \xrightarrow{\text{Time 1}} P = \begin{bmatrix} .8 & .1 & .1 \\ .1 & .9 & 0 \\ .1 & 0 & .9 \end{bmatrix} \begin{bmatrix} .2 \\ .3 \\ .5 \end{bmatrix} = \begin{bmatrix} .24 \\ .29 \\ .47 \end{bmatrix} \xrightarrow{\text{Time 2}} P = \begin{bmatrix} .8 & .1 & .1 \\ .1 & .9 & 0 \\ .1 & 0 & .9 \end{bmatrix} \begin{bmatrix} .24 \\ .29 \\ .47 \end{bmatrix} = \begin{bmatrix} .268 \\ .285 \\ .447 \end{bmatrix}$$

$$\text{Time } n \xrightarrow{\quad} P = \begin{bmatrix} .8 & .1 & .1 \\ .1 & .9 & 0 \\ .1 & 0 & .9 \end{bmatrix}^n \begin{bmatrix} .2 \\ .3 \\ .5 \end{bmatrix}$$

# Chapman–Kolmogorov Equations

9

- Generalization of how the distribution of states of a Markov model evolves over time
- Suppose we have a Markov model with  $|Q|$  states where we define  $p_{ij}(n)$  to be the probability of going from state  $i$  to state  $j$  in exactly  $n$  steps

$$p_{ij}(n + m) = \sum_{k=1}^{|Q|} p_{ik}(n) p_{kj}(m)$$

for all  $n \geq 0$ ,  $m \geq 0$ , and any states  $i$  and  $j$ .

That is, the probability of getting to state  $j$  from state  $i$  in  $(n+m)$  steps is the sum over all possible intermediate states  $k$  of the probability of getting from  $i$  to  $k$  in  $n$  steps, then from  $k$  to  $j$  in the remaining  $m$  steps

# Stationary Distributions

10

- Look at the evolution of Markov model over really long time scale for previous example:

$$\begin{bmatrix} .2 \\ .3 \\ .5 \end{bmatrix} \rightarrow \begin{bmatrix} .24 \\ .29 \\ .47 \end{bmatrix} \rightarrow \begin{bmatrix} .268 \\ .285 \\ .447 \end{bmatrix} \rightarrow \dots \rightarrow \begin{bmatrix} .33333 \\ .33333 \\ .33333 \end{bmatrix} \rightarrow \begin{bmatrix} .33333 \\ .33333 \\ .33333 \end{bmatrix}$$

- Convergence on a single probability distribution that will not change on further multiplication
- Always converge to the same final distribution vector, regardless of our starting point (initial distribution)

This vector on which the state distribution converges after a large number of steps is called the **stationary distribution**

# Stationary Distributions

11

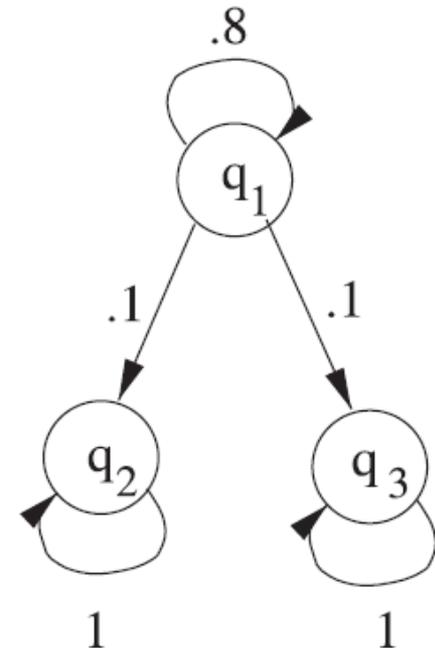
- Will this property of convergence on a unique stationary distribution regardless of starting point work for any example?

- ▣ Answer is NO. It is possible that final vector is not unique!

- Example:

$$P = \begin{bmatrix} .8 & 0 & 0 \\ .1 & 1 & 0 \\ .1 & 0 & 1 \end{bmatrix}$$

Start	Final
$\begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix}$	$\begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix}$
$\begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$	$\begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$



# Stationary Distributions

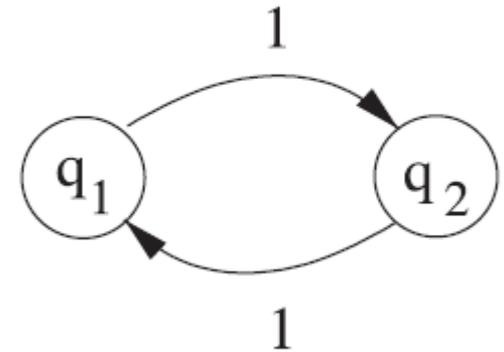
12

- A Markov model is not even guaranteed to converge on any vector
- Example:

$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

Initial Probability

$$\begin{bmatrix} p \\ 1 - p \end{bmatrix}$$



$$\begin{bmatrix} p \\ 1 - p \end{bmatrix} \rightarrow \begin{bmatrix} 1 - p \\ p \end{bmatrix} \rightarrow \begin{bmatrix} p \\ 1 - p \end{bmatrix} \rightarrow \begin{bmatrix} 1 - p \\ p \end{bmatrix} \rightarrow \begin{bmatrix} p \\ 1 - p \end{bmatrix} \rightarrow \dots$$

# Ergodicity

13

- Ergodicity means that for any two states  $q_i$  and  $q_j$  there is some sequence of transitions with nonzero probability that go from  $q_i$  to  $q_j$ 
  - ▣ Ergodic Markov chain is also sometimes called irreducible

## □ Example 1

$$\begin{bmatrix} .8 & .1 & .1 \\ .1 & .9 & 0 \\ .1 & 0 & .9 \end{bmatrix}$$

Ergodic



Unique

Stationary Distribution

## Example 2

$$\begin{bmatrix} .8 & 0 & 0 \\ .1 & 1 & 0 \\ .1 & 0 & 1 \end{bmatrix}$$

Not Ergodic



No Unique

Stationary Distribution

## Example 3

$$\begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

Not Ergodic



No

Stationary Distribution

# Eigenvalues and Stationary Distribution

14

- Markov model will converge to a unique stationary distribution if its transition matrix has exactly one eigenvector with eigenvalue  $\lambda_1=1$  and has  $|\lambda_i| < 1$  for every other eigenvector
  - Similar to Power Method of computing maximum eigenvalue and its corresponding eigenvector
  - Converges to this eigenvector after all eigenvalues die out after k-iterations:  $\lambda_i^k = 1$  ( $i=1$ ) or 0 (otherwise)

# Eigenvalues and Stationary Distribution

15

- If a Markov model is not ergodic, then its state set can be partitioned into discrete graph components unreachable from one another
  - ▣ Each such component will have its own eigenvector with eigenvalue 1
  - ▣ Depending on which component we start in, we may converge on any of them
- Example: Nonergodic Markov model 2

$$\begin{bmatrix} .8 & 0 & 0 \\ .1 & 1 & 0 \\ .1 & 0 & 1 \end{bmatrix} \rightarrow \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix} \quad \text{and} \quad \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix} \quad \text{Are both eigenvector with eigenvalue}=1$$

# Test of Markov Model Convergence

16

- A Markov model is guaranteed to converge on a stationary distribution if there exists some integer  $N > 0$  such that,

$$\min_{i,j} p_{ij}(N) = \delta, \quad \delta > 0.$$

- That is, there is some number of steps  $N$  such that no matter where we start, we have some bounded nonzero probability of getting to any given ending position in exactly  $N$  steps.

# Mixing Time

17

- Informally, the mixing time is the time needed for the Markov model to get close to its stationary distribution
- if we want to run the model long enough for the transients to die away by some factor  $r$ , then we need to run for a number of rounds  $k$  such that,
  - ▣ Assume  $\lambda_1=1$  and  $|\lambda_i| < 1, i \neq 1$

$$|\lambda_2|^k = r$$

$$k \log \lambda_2 = \log r$$

$$k = \frac{\log r}{\log \lambda_2}.$$

# Assignments

18

- For each of the following models:
  - ▣ Determine whether the following Markov models have stationary distributions
  - ▣ Estimate stationary distributions (if available)
  - ▣ Compare Stationary distributions to eigenvector corresponding to maximum eigenvalue (if available)
  - ▣ Estimate mixing time for transients to die out by a factor of  $1/100$

$$P_1 = \begin{bmatrix} 1 & 0.2 & 0.1 \\ 0 & 0.6 & 0.2 \\ 0 & 0.2 & 0.7 \end{bmatrix}$$

$$P_3 = \begin{bmatrix} 0.6 & 0.4 \\ 0.4 & 0.6 \end{bmatrix}$$

Assume initial state of  $q_1$   
for all models

$$P_2 = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 0.8 & 0.3 \\ 0 & 0.2 & 0.7 \end{bmatrix}$$

$$P_4 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$